Lecture 7: Computation of Greeks

Ahmed Kebaier kebaier@math.univ-paris13.fr

HEC, Paris







Motivation

• The pathwise method requires some restrictive regularity assumptions on the payoff function of the option price, at least one time differentiable. This is not the case of some particular payoffs such as the digital options with a payoff function

$$h_T = h(X_T^x)$$
 where $h(x) = \mathbf{1}_{\{x \ge K\}}$.

• The Greek Δ of such option cannot be evaluated using the pathwise method.

• Note also that for the same arguments we cannot use the pathwise method to evaluate the Greek payoff Γ of a given vanilla Call option.

• Let us assume that the family of random variables $(X(\theta))$ indexed by a parameter $\theta \in \Theta$ (Θ an open set of \mathbb{R}), admits positive density function $p(\theta, y)$.

< □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > <

• Let us assume that the family of random variables $(X(\theta))$ indexed by a parameter $\theta \in \Theta$ (Θ an open set of \mathbb{R}), admits positive density function $p(\theta, y)$.

$$f(heta) = \mathbb{E}\phi(X(heta)) = \int_{\mathbb{R}^d} \phi(y) p(heta, y) dy.$$

< □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > <

• Let us assume that the family of random variables $(X(\theta))$ indexed by a parameter $\theta \in \Theta$ (Θ an open set of \mathbb{R}), admits positive density function $p(\theta, y)$.

$$f(heta) = \mathbb{E}\phi(X(heta)) = \int_{\mathbb{R}^d} \phi(y) p(heta, y) dy.$$

Theorem 1

Assume that the density function $p(\theta, y)$ taking values in $r \Theta \times \mathbb{R}^d$ satisfies

i) $\theta \mapsto p(\theta, y)$ is differentiable on Θ almost every where

• Let us assume that the family of random variables $(X(\theta))$ indexed by a parameter $\theta \in \Theta$ (Θ an open set of \mathbb{R}), admits positive density function $p(\theta, y)$.

$$f(\theta) = \mathbb{E}\phi(X(\theta)) = \int_{\mathbb{R}^d} \phi(y) p(\theta, y) dy.$$

Theorem 1

Assume that the density function $p(\theta, y)$ taking values in $r \Theta \times \mathbb{R}^d$ satisfies

i) $\theta \mapsto p(\theta, y)$ is differentiable on Θ almost every where ii) $\exists g : \mathbb{R}^d \mapsto \mathbb{R}$ a measurable function such that $\int_{\mathbb{R}^d} \phi(y)g(y)dy < \infty$ and $\forall \theta \in \Theta \ |\partial_{\theta}p(\theta, y)| \le g(y)$,

• Let us assume that the family of random variables $(X(\theta))$ indexed by a parameter $\theta \in \Theta$ (Θ an open set of \mathbb{R}), admits positive density function $p(\theta, y)$.

$$f(\theta) = \mathbb{E}\phi(X(\theta)) = \int_{\mathbb{R}^d} \phi(y) p(\theta, y) dy.$$

Theorem 1

Assume that the density function $p(\theta, y)$ taking values in $r \Theta \times \mathbb{R}^d$ satisfies

i) $\theta \mapsto p(\theta, y)$ is differentiable on Θ almost every where ii) $\exists g : \mathbb{R}^d \mapsto \mathbb{R}$ a measurable function such that $\int_{\mathbb{R}^d} \phi(y)g(y)dy < \infty$ and $\forall \theta \in \Theta \ |\partial_{\theta}p(\theta, y)| \le g(y)$, Then

$$orall heta \in \Theta, \quad f'(heta) = \mathbb{E}\left(\phi(X(heta)) rac{\partial \log p}{\partial heta}(heta, X(heta))
ight)$$

Proof.

• According to assumptions i) and ii) we have

$$f'(heta) = rac{\partial}{\partial heta} \mathbb{E} \phi(X(heta)) = \int_{\mathbb{R}^d} \phi(y) rac{\partial}{\partial heta} p(heta, y) dy.$$

• Rewriting the above expression we get

$$f'(\theta) = \int_{\mathbb{R}^d} \phi(y) \frac{\frac{\partial}{\partial \theta} p(\theta, y)}{p(\theta, y)} p(\theta, y) dy = \mathbb{E} \left(\phi(X(\theta)) \frac{\frac{\partial}{\partial \theta} p(\theta, X(\theta))}{p(\theta, X(\theta))} \right),$$

◆□▶ ◆□▶ ◆三▶ ◆三▶ 三三 のへぐ

which completes the proof.

• Recall that the Black-Scholes model with parameters r, σ, T is given by the explicit solution

$$S_T = S_0 \exp\left(\left(r - \sigma^2/2\right)T + \sigma\sqrt{T}G
ight), \quad ext{ where } \quad G \sim \mathcal{N}(0,1).$$

• Consequently, for any given measurable function we get

$$\mathbb{E}f(S_T) = \mathbb{E}f\left(S_0 \exp\left((r - \sigma^2/2)T + \sigma\sqrt{T}G\right)\right) \\ = \int_{\mathbb{R}} f\left(S_0 \exp\left((r - \sigma^2/2)T + \sigma\sqrt{T}y\right)\right) \\ \times \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{y^2}{2}\right) dy.$$

• Using the following change of variable

$$u = S_0 \exp\left((r - \sigma^2/2)T + \sigma\sqrt{T}y\right) \text{ with}$$
$$du = \sigma\sqrt{T}S_0 \exp\left((r - \sigma^2/2)T + \sigma\sqrt{T}y\right)dy = u\sigma\sqrt{T}dy$$
and $y = \zeta(u)$ where $\zeta(u) = \left(\log(u/S_0) - (r - \sigma^2/2)T\right)/\sigma\sqrt{T}$

• Using the following change of variable

$$u = S_0 \exp\left((r - \sigma^2/2)T + \sigma\sqrt{T}y\right) \quad \text{with}$$
$$du = \sigma\sqrt{T}S_0 \exp\left((r - \sigma^2/2)T + \sigma\sqrt{T}y\right)dy = u\sigma\sqrt{T}dy$$
and $y = \zeta(u)$ where $\zeta(u) = \left(\log(u/S_0) - (r - \sigma^2/2)T\right)/\sigma\sqrt{T}$

we get

$$\mathbb{E}f(S_{T}) = \int_{\mathbb{R}} f(y) \frac{1}{u\sigma\sqrt{T}\sqrt{2\pi}} \exp\left(-\frac{\zeta(u)^{2}}{2}\right) du.$$

• Therefore we deduce that the density of the random variable S_T is given by

$$g(x) = rac{1}{x\sigma\sqrt{T}\sqrt{2\pi}}\exp\left(-rac{\zeta(x)^2}{2}
ight).$$

• Therefore we deduce that the density of the random variable S_T is given by

$$g(x) = rac{1}{x\sigma\sqrt{T}\sqrt{2\pi}}\exp\left(-rac{\zeta(x)^2}{2}
ight).$$

Then

$$\frac{\partial \log g(x)}{\partial S_0} = \frac{\frac{\partial g}{\partial S_0}(x)}{g(x)} = \frac{\log(x/S_0) - (r - \sigma^2/2)T}{S_0 \sigma^2 T}.$$

◆□▶ ◆□▶ ◆三▶ ◆三▶ 三三 のへぐ

Hence, according to the above theorem

$$\begin{aligned} \Delta &= \frac{\partial}{\partial S_0} e^{-rT} \mathbb{E} (S_T - K)_+ \\ &= e^{-rT} \mathbb{E} \left((S_T - K)_+ \frac{\log(S_T/S_0) - (r - \sigma^2/2)T}{S_0 \sigma^2 T} \right) \\ &= e^{-rT} \mathbb{E} \left((S_T - K)_+ \frac{G}{S_0 \sigma \sqrt{T}} \right). \end{aligned}$$

▲□▶ ▲圖▶ ▲≣▶ ▲≣▶ ▲国 ● ● ●

• Let us consider the case of an Asian option. Since the associated payoff involves the vector $(S_{t_1}, S_{t_2}, \cdots, S_{t_n})$

< □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > <

• Let us consider the case of an Asian option. Since the associated payoff involves the vector $(S_{t_1}, S_{t_2}, \cdots, S_{t_n})$

• we need to make explicit the joint density of this vector if we aim to apply the log-likelihood ratio method for the computation of sensitivities associated to the Asian option price.

• Let us consider the case of an Asian option. Since the associated payoff involves the vector $(S_{t_1}, S_{t_2}, \cdots, S_{t_n})$

• we need to make explicit the joint density of this vector if we aim to apply the log-likelihood ratio method for the computation of sensitivities associated to the Asian option price.

• Using the Markovian property of the Brownian motion we can rewrite the density associated to the above vector as follows

$$g(x_1, \cdots, x_n) = g_1(x_1|S_0)g_2(x_2|x_1)\cdots g_n(x_n|x_{n-1}),$$
 with

$$g_j(x_j|x_{j-1}) = \frac{1}{x_i \sigma \sqrt{t_j - t_{j-1}}} \varphi(\zeta_j(x_j|x_{j-1})), \quad \text{where}$$
$$\varphi(u) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{u^2}{2}\right)$$

$$\zeta_j(x_j|x_{j-1}) = \frac{\log(x_j/x_{j-1}) - (r - \sigma^2/2)(t_j - t_{j-1})}{\sigma\sqrt{t_j - t_{j-1}}}.$$

$$\zeta_j(x_j|x_{j-1}) = \frac{\log(x_j/x_{j-1}) - (r - \sigma^2/2)(t_j - t_{j-1})}{\sigma\sqrt{t_j - t_{j-1}}}.$$

Consequently,

$$\frac{\partial \log g}{\partial S_0}(S_{t_1}, \cdots, S_{t_n}) = \frac{\partial \log g_1}{\partial S_0}(S_{t_1}|S_0) = \frac{G_1}{S_0 \sigma \sqrt{t_1}},$$

◆□▶ ◆□▶ ◆三▶ ◆三▶ 三三 のへぐ

where G_1 is the same Gaussian used for the simulation of S_{t_1} .

$$\zeta_j(x_j|x_{j-1}) = \frac{\log(x_j/x_{j-1}) - (r - \sigma^2/2)(t_j - t_{j-1})}{\sigma\sqrt{t_j - t_{j-1}}}.$$

Consequently,

$$\frac{\partial \log g}{\partial S_0}(S_{t_1}, \cdots, S_{t_n}) = \frac{\partial \log g_1}{\partial S_0}(S_{t_1}|S_0) = \frac{G_1}{S_0 \sigma \sqrt{t_1}},$$

where G_1 is the same Gaussian used for the simulation of S_{t_1} .

• Finally, the value of the delta Asian call option using the Log-likelihood ratio method is given by

$$\Delta = e^{-rT} \mathbb{E}^* \left(\left(\frac{1}{n} \sum_{i=1}^n S_{t_i} - \mathcal{K} \right)_+ \frac{G_1}{S_0 \sigma \sqrt{t_1}} \right).$$

◆□▶ ◆□▶ ◆臣▶ ◆臣▶ 三臣 - のへで